Efficacy of platelet-rich plasma injections for symptomatic tendinopathy: systematic review and meta-analysis of randomised injection-controlled trials

Larry E Miller,¹ William R Parrish,² Breana Roides,² Samir Bhattacharyya²

ABSTRACT

Aim To determine the efficacy of platelet-rich plasma (PRP) injections for symptomatic tendinopathy.

Design Systematic review of randomised, injection-controlled trials with meta-analysis.

Data sources Systematic searches of MEDLINE and EMBASE, supplemented by manual searches.

Eligibility criteria for selecting studies Randomised controlled trials with 3 months minimum follow-up that evaluated pain reduction with PRP versus control (saline, local anaesthetic, corticosteroid) injections in patients with symptomatic tendinopathy.

Results A total of 16 randomised controlled trials (18 groups) of PRP versus control were included. Median sample size was 35 patients, a study size that would require an effect size ≥1.0 to achieve statistical significance. PRP was more efficacious than control in reducing tendinopathy pain, with an effect size of 0.47 (95% CI 0.22 to 0.72, p<0.001), signifying a moderate treatment effect. Heterogeneity among studies was moderate (I²=67%, p<0.001). In subgroup analysis and meta-regression, studies with a higher proportion of female patients were associated with greater treatment benefits with PRP.

Conclusions Injection of PRP is more efficacious than control injections in patients with symptomatic tendinopathy.

INTRODUCTION

Tendinopathy is the most common musculoskeletal complaint in patients seeking medical care. The most common sites of presentation include the elbow, rotator cuff, Achilles tendon and patellar tendon. With early diagnosis and timely application of traditional non-surgical treatments such as activity modification, gentle static stretching, anti-inflammatory medications and/or eccentric loading, the prognosis is favourable in the acute stage. However, symptoms may persist in some patients despite exhausting these treatment options. Recalcitrant tendinopathy may manifest because, once damaged, the biological and biomechanical properties of connective tissue are never completely restored. Healing times in chronic tendinopathies are prolonged as tendons are relatively hypovascular and local blood flow is only about one-third of that delivered to the muscles.

Since the aetiology of chronic tendinopathy is controversial and likely multifactorial, numerous therapies with various mechanisms of action have been attempted although none have an ideal efficacy and safety profile. Therapies intended to reduce inflammation such as local cooling, non-steroidal anti-inflammatory drugs and corticosteroids are commonly prescribed, yet the premise for application of these modalities is misguided given the...
absence of a measurable inflammatory response in or around the lesion in chronic tendinopathy. This may partially explain their limited efficacy in recalcitrant tendinopathy cases.

Tendinosis forms as an imbalance between the demands that are placed on a tendon and its ability to remodel. Recent developments in biological research have emphasised the importance of growth factors in the maintenance of normal tissue structure and repair of tissue lesions. Platelet-rich plasma (PRP) is a blood derivative with a platelet concentration greater than that of whole blood that is an emerging regenerative therapy for tissue injury and degeneration. Activated platelets release biologically active proteins that promote cellular recruitment, growth and morphogenesis. Soft tissue healing is thought to be stimulated via enhanced fibroblast migration and proliferation, upregulated vascularisation and increased collagen deposition.

These biological properties are appealing in the treatment of tendinopathy, which has poor intrinsic healing ability. Among the studies performed on PRP injection in the treatment of tendinopathy, study designs and PRP preparation methods have varied widely which complicates interpretation of PRP efficacy. The purpose of this systematic review and meta-analysis was to determine the efficacy of PRP injections for tendinopathy by evaluating randomised controlled trials of PRP injection versus control injection. A secondary purpose of this research was to explore sources of heterogeneity in treatment outcomes among studies.

METHODS

Study selection

The study was performed according to PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-analyses). Two researchers independently searched MEDLINE and EMBASE for randomised controlled trials of PRP injection versus control injections (saline, local anaesthetic or corticosteroid) for treatment of tendinopathy using a combination of diagnostic and therapy-specific keywords and MeSH terms. The details of the MEDLINE search strategy are listed in box 1. The syntax for EMBASE was similar, but adapted as necessary. Additionally, reference lists of included papers and relevant meta-analyses were manually searched. No date or language restrictions were applied to the searches. The final search was performed on 30 November 2016. Main inclusion criteria included randomised controlled trial of PRP injection; control group treated with control injection (saline, local anaesthetic or corticosteroid); primary diagnosis of symptomatic tendinopathy; minimum follow-up period of 3 months; and extractable measures of pain at baseline and post-treatment. When multiple studies included overlapping series of patients, only the study with the largest sample size or longest follow-up duration was included. Study selection discrepancies between the two researchers were resolved by discussion.

Data extraction and quality assessment

An initial database was developed, pilot tested and refined to ensure consistency with outcomes reported in the literature. Data were extracted from eligible peer-reviewed articles by one researcher and verified by a second researcher; data extraction discrepancies were resolved by discussion. The Cochrane Collaboration tool was used for assessing risk of bias in randomised trials. The risk of bias tool assesses sequence generation, allocation concealment, blinding, incomplete outcome data, selective outcome reporting and other sources of bias. Assessments of the risk of bias were categorised as high, low or uncertain for each item in a given study.

Outcomes

Tendinopathy pain severity was the efficacy outcome of interest in this analysis. Pain severity on a visual analogue scale (VAS) was preferentially extracted from each study. When not reported, data were extracted from relevant pain severity tools reported in each study. Data from the final follow-up period between 3 and 12 months were used in the main analysis.

Data analysis

A random effects meta-analysis model was selected a priori for all analyses. The effect size was reported as the
standard mean difference (SMD) for PRP relative to control injection, respectively. For reference, SMD values of 0.2, 0.5, 0.8 and 1.0 are defined as small, medium, large and very large effect sizes, respectively. When a single PRP group was compared with multiple control groups within a study, the sample size of the PRP group entered into the meta-analysis was adjusted based on the number of control groups. Forest plots were used to visually assess effect sizes and corresponding 95% CIs across studies. Publication bias was visually assessed with a funnel plot and quantitatively assessed using Egger’s regression test. The I² statistic was used to estimate heterogeneity of treatment effects among studies with values of ≤25%, 50% and ≥75% representing low, moderate and high inconsistency, respectively. Post hoc subgroup analyses and meta-regression were undertaken to explore sources of heterogeneity among studies in pain severity. A one-study removed sensitivity analysis was performed, which recalculates the meta-analysis after removing one study at a time in order to explore the impact of single studies on treatment effects. p Values were two sided with a significance level <0.05. All analyses were performed using Comprehensive Meta-analysis (V.2.2, Biostat, Englewood, NJ, USA).

### RESULTS

#### Study selection

After screening 626 records for eligibility, 16 randomised controlled trials (18 groups) of PRP versus control injections were included. The most common reasons for exclusion were review paper (53), non-tendinopathy (14), injection of active control (7; eg, whole blood, PRP) and non-randomised study (7). Several manuscripts were excluded from this review because they were superseded by papers from the same study with longer follow-up including de Vos et al. and Peerbooms et al. A flow diagram of study identification and selection is shown in figure 1.

#### Patient and study characteristics

Baseline patient characteristics were comparable between the PRP and control groups (table 1). Overall, 54% of patients were female and median age was 48 years. Minimum tendinopathy symptom duration ranged from 1 to 6 months (unreported in two studies). Median sample size was 35 patients, a study size that would require an effect size ≥1.0 to achieve statistical significance. Maximum follow-up ranged from 3 to 24 months. The PRP preparation methods and injection protocols used in each study are detailed in table 2. Single injection protocols were used in 81% of studies. The most common PRP characteristics were ≥5× platelet concentration (10 of 18 groups), increased leukocyte concentration (9 of 17 groups; 1 group unreported) and no PRP activation (16 of 17 groups; 1 group unreported). Controls consisted of saline and/or anaesthetic injection in 11 groups and corticosteroid with or without anaesthetic injection in 7 groups. Risk of bias assessment for each study is detailed in table 3. No study was determined to have low risk of bias, 5 studies had uncertain risk of bias and 11 were at high risk of bias.

![Figure 1 PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) flow diagram.](image-url)
<table>
<thead>
<tr>
<th>Study</th>
<th>Treatment period</th>
<th>Sample size</th>
<th>Female gender (%)</th>
<th>Age (year)</th>
<th>Tendinopathy</th>
<th>Minimum symptom duration (month)</th>
<th>Pain outcome</th>
<th>Follow-up (month)</th>
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— indicates missing data.

*Patients failed to respond to ≥6 weeks physical therapy; total symptom duration not reported.

†Data extracted through 12 months only for meta-analysis per systematic review methods (ie, data extraction at 3, 6 and 12 months); however, total symptom duration not reported.

‡Study includes same PRP group and different control groups. PRP group sample size adjusted in meta-analysis based on number of groups.

§Maximum symptom duration was 3 months.

¶Data extracted through 3 months only for meta-analysis due to implausible reported 6-month outcomes.

ASES, American Shoulder and Elbow Surgeons; DASH, Disabilities of the Arm, Shoulder and Hand score; PRP, platelet-rich plasma; PRTEE, Patient-Rated Tennis Elbow Evaluation; SPADI, Shoulder Pain and Disability Index; VAS, visual analogue scale; VISA-A, Victorian Institute of Sport Assessment-Achilles questionnaire.
PRP efficacy

Injection of PRP resulted in statistically lower pain severity relative to control in 8 of 18 groups, 10 of 18 groups reported no differences, and no groups reported greater efficacy with control over PRP. In the random effects meta-analysis, PRP was associated with lower tendinopathy pain severity. The SMD for PRP was 0.47 (95% CI 0.22 to 0.72, p<0.001), which is considered a moderate treatment effect (figure 2). No evidence of publication bias (Egger’s regression p=0.66; figure 3) was found. Heterogeneity among studies was moderate ($I^2=67\%$, p<0.001). Potential sources of heterogeneity in treatment effects were investigated with subgroup analyses (table 4). The only variable that was shown to influence PRP efficacy was female sex. In the eight groups with proportion of women above the overall median, a large treatment effect was observed (SMD=0.71). However, in the remaining six groups with fewer women, the treatment benefit was negligible (SMD=0.11). The relationship between female sex and PRP efficacy was further confirmed in meta-regression where the proportion of women in each study explained 34% of the variability in treatment effects (p<0.001) (figure 4). In the subgroup analysis, other variables that

<table>
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<th>Study</th>
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<th>Injection contents and volume†</th>
<th>Control</th>
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<td>Behera et al</td>
<td>4B</td>
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<td>3 mL bupivacaine, 0.5 mL normal saline</td>
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<td>2§</td>
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†From Mishra et al (1A=high platelet concentration with leucocyte counts > whole blood and no exogenous platelet activation, 3B=low platelet concentration with leucocyte counts < whole blood and no exogenous platelet activation, 4B=3B but with exogenous platelet activation).

§Volumen unspecified.

¶Injections separated by 4-week interval.

PRP, platelet-rich plasma.
were associated with clinically meaningful differences in PRP efficacy (SMD ≥ 0.3), yet did not achieve statistical significance, were tendinopathy location (lateral epicondylar over other locations) and maximum follow-up duration (12 months over 3 months). Thus, while these comparisons were underpowered, the results suggest that PRP may have greater efficacy in lateral epicondylar tendinopathy or with longer follow-up duration. The ‘one study removed’ sensitivity analysis demonstrated that no single study significantly altered conclusions of the main analyses when removed from the analysis, with the SMD in all scenarios ranging from 0.39 to 0.51 (all p ≤ 0.001) (figure 5).

**DISCUSSION**

The results of this systematic review and meta-analysis provide level 1 evidence that injection of PRP is efficacious in patients with symptomatic tendinopathy. The treatment effects with PRP relative to controls in this meta-analysis suggest clinically meaningful improvements in patient symptoms. Previous meta-analyses have drawn disparate conclusions regarding PRP efficacy, likely because of widely varying methodologies among studies.32–38 We designed the current review to minimise potential sources of bias, namely by excluding non-randomised studies, studies with non-injection control groups, or...
active injectable controls (eg, whole blood, PRP). Still, we identified significant heterogeneity in treatment effects. When evaluating patient, treatment and study design-related factors, female sex was the only variable that modified the efficacy of PRP for treatment of tendinopathy. The observation that PRP may be more efficacious in women has been previously reported. Wesner et al reported that the magnitude of pain reduction on a 0–10 scale was greater in women than men (2.8 vs 1.8, p=0.04) with PRP injection in degenerative tendinopathy. While no obvious explanation exists for this post hoc observation, exploration of gender differences with PRP injection should be explored in future studies.

Lateral epicondylar tendinopathy was evaluated in most comparisons (12 of 18 groups) and was the most responsive to PRP therapy (effect size=0.57). For comparison, rotator cuff (three groups; effect size=0.32), Achilles (two groups; effect size=0.22) and patellar tendon (one group; effect size=−0.13) pathology were less studied and had negligible to small effect sizes. In agreement with our findings, others have reported that PRP is particularly efficacious for lateral epicondylar tendinopathy. While this meta-analysis was underpowered to detect meaningful differences in treatment effects among anatomical sites, it is plausible that PRP efficacy may also be influenced by injection site.

**Figure 2** Forest plot of platelet-rich plasma (PRP) versus control on tendinopathy pain. Random effects meta-analysis using the standard mean difference statistic for PRP versus control. A pooled estimate of overall standard mean difference (diamond) and 95% CI (diamond width) summarises the effect size. Standard mean difference values of 0.2, 0.5, 0.8 and 1.0 are defined as small, medium, large and very large effect sizes, respectively. Effects to the left of 0 indicate greater efficacy with control; effects to the right of 0 indicate greater efficacy with PRP. When the horizontal bars of an individual study, or the pooled diamond width, cross 0, the effect is not significantly different. Heterogeneity: $I^2=67\%$, p<0.001. SMD, standard mean difference.

**Figure 3** Funnel plot of standard mean difference in platelet-rich plasma efficacy across studies. Egger’s p value=0.66 for publication bias. SMD, standard mean difference.
Two randomised controlled trials that were included in this meta-analysis warrant additional discussion. First, in the study of Behera and colleagues, the treatment benefit of PRP relative to control was considerably greater than any other included study (effect size = 2.2). Over 1-year follow-up, pain scores on a 0–100 scale decreased from 75 ± 6 to 13 ± 14 with PRP and from 76 ± 7 to 41 ± 12 with control (bupivacaine) injection. While exclusion of this study in a one-study removed analysis did not change the conclusions of this meta-analysis, heterogeneity in outcomes among the remaining studies was non-existent following removal of this study. Although no specific attributes of this study that may dramatically impact outcomes are readily observable, the inclusion of this study does introduce considerable inconsistency to our findings.

Second, the study by Dragoo and colleagues reported an unprecedented and profound recovery in the bupivacaine control group at the final time point. For example, VAS pain scores in the control group were 3.0 ± 2.3 at baseline, 2.3 ± 1.6 at 12 weeks and 0.3 ± 0.5 at 26 weeks. This 90% reduction in mean pain severity with control over a 6-month time frame was notably greater than in any other study of control injections. Complete recovery of a non-active control group in a randomised controlled trial for an orthopaedics indication is unanticipated. As before, exclusion of this study in a one-study removed analysis did not change the conclusions of this meta-analysis and no specific attributes of this study that may dramatically impact outcomes are readily observable.

The sample sizes of most PRP studies for symptomatic tendinopathy were too small to statistically detect clinically meaningful treatment effects. To detect the effect size of PRP observed in this meta-analysis, a sample size of 146 patients (73 per group) would be required. In this review, only 1 of 18 groups enrolled at least this number of patients. In fact, the median sample size in this review was only 35 patients. This observation likely explains why many individual studies showed no benefit of PRP, yet the results of the pooled analysis showed a statistically significant, moderate benefit relative to control injections.
A similar observation can be made about the systematic review by de Vos et al. who concluded that there was strong evidence against PRP for chronic lateral epicondylar tendinopathy. This conclusion was based on the observation that only one of six included studies showed a positive benefit of PRP, yet no attempt was made at quantitative data synthesis. (What was the effect size and direction of the analysis when you reviewed it?) These results underscore the need for investigators to perform power analyses with realistic assumptions during study planning and for systematic reviewers to consider meta-analytic techniques, where appropriate, to quantify treatment effects with more precision than simple counts of positive studies.

Our meta-analysis is associated with several issues that may influence interpretation. Strengths of this meta-analysis are inclusion of only randomised, injection-controlled trials, structured data extraction methodology and comprehensive analysis of potentially confounding factors. There were also limitations inherent in the studies that were included in this review. First, the duration of tendinopathy symptoms was variable, frequently of short duration, and, in many cases, inadequately described. Thus, this meta-analysis was unable to discern the efficacy of PRP based on chronology of symptoms. Second, there was significant heterogeneity in efficacy outcomes among studies with PRP versus control injections. While subgroup and meta-regression identified female sex as a potential mediating factor, definitive conclusions cannot be drawn given the post hoc nature of the analysis. Third, the duration of patient follow-up may be an important

![Figure 4](image)

**Figure 4** Meta-regression of relationship between proportion of women in each study and platelet-rich plasma efficacy. Percentage of explained variance=34%, p<0.001. Markers are proportional to sample size. SMD, standard mean difference.

![Figure 5](image)

**Figure 5** Forest plot for one-study removed sensitivity analysis of platelet-rich plasma (PRP) versus control on tendinopathy pain. Random effects meta-analysis using the standard mean difference statistic for PRP versus control. A pooled estimate of overall standard mean difference (diamond) and 95% CI (diamond width) summarises the effect size. Standard mean difference values of 0.2, 0.5, 0.8 and 1.0 are defined as small, medium, large and very large effect sizes, respectively. Effects to the left of 0 indicate greater efficacy with control; effects to the right of 0 indicate greater efficacy with PRP. When the horizontal bars of an individual study, or the pooled diamond width, cross 0, the effect is not significantly different. SMD, standard mean difference.

determinant of PRP efficacy. Subgroup analyses demonstrated greater effect sizes with PRP with greater follow-up duration. Although the analysis was under-powered to detect important differences given the limited number of studies with varying follow-up durations, the magnitude of the effect size with PRP at 12 months and the effect size difference from 3 to 12 months suggests that PRP efficacy may continue to improve through at least 12 months follow-up. Thus, researchers are encouraged to enrol an adequate number of patients and continue follow-up through at least 1 year post-treatment. Third, we made no attempt to assess safety of PRP injections in this study. Generally, safety reporting in the PRP literature is inconsistent and inadequate. While treatment-related complications with PRP such as pain and swelling are generally infrequent, mild and transient, the potential for unreported complications remains a major limitation of the PRP literature in general and the consequent absence of pooled safety data is a limitation of this meta-analysis.

CONCLUSIONS
Injection of PRP is more efficacious than control injections in patients with symptomatic tendinopathy. Indications for use and PRP preparation methods should continue to be refined in an effort to reduce variability in outcomes and identify optimal treatment conditions.

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REFERENCES


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Larry E Miller, William R Parrish, Breana Roides and Samir Bhattacharyya


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